Validly Estimating True Dose-Response When Only Treatment versus Control is Randomized: Principal Stratification for Causal Inference with Extended Partial Compliance

Hui Jin & Donald B. Rubin

Department of Statistics Harvard University

September, 2006



Overview

Background: Efron and Feldman (1991) - EF:

- One of the earliest statistical articles to address non-compliance in randomized experiments.
- EF analyzed data from the Lipid Research Clinics Coronary Primary Prevention Trial (LRC-CPPT) to study the effectiveness of cholestyramine for lowering cholesterol levels.
- LRC-CPPT: Randomized treatment versus placebo, not dose.
- EF discussed inference for "dose-response" from non-randomized data.

Overview

Our work:

- Analyze the same data within the framework of Principal Stratification (Frangakis and Rubin, 2002).
- Explicate possible assumptions, including more flexible ones.
- Check EF's assumptions within our model.
- Formalize inference for dose-response.
- Our idea applies to any setting where dose is not randomized, e.g., amount of studying, hours of job-training.

LRC-CPPT Data

Specific features of LRC-CPPT:

- Placebo-controlled double blind randomized clinical trial to study the effectiveness of cholestyramine.
- 164 men were randomized to the treatment group and assigned the drug.
- 171 men were randomized to the control group and assigned placebo.
- For each patient, cholesterol levels were measured before and after taking the drug (or placebo).
- The outcome variable, Y, was the decrease in cholesterol level: the only variable available to EF or to us, besides treatment assigned and dose taken.



LRC-CPPT Data

Partial Compliance Complications:

- Most patients in the treatment group only took a proportion of the assigned drug.
- Most patients in the control group only took a proportion of the assigned placebo.

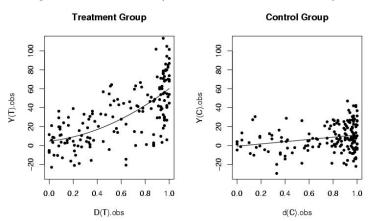
Data Available:

- Z_i: treatment assignment
- D_i(T) or d_i(C): compliance to drug under treatment or compliance to placebo under control
- Y_i(T) or Y_i(C): outcome under treatment or outcome under control



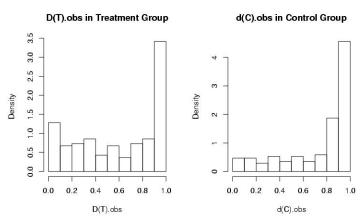
LRC-CPPT Data: Figure 1

Figure: Observed Compliance-Outcome Relationship



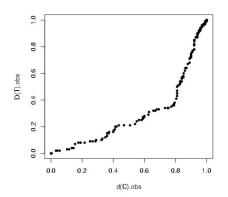
LRC-CPPT Data: Figure 2

Figure: Histograms of Observed Compliances



LRC-CPPT Data: Figure 3

Figure: Q-Q Plot of Observed Drug and Placebo Compliances



Full Principal Stratification with Extended Compliance

| i | Xi | Z_i | $D_i(T)$ | $D_i(C)$ | $d_i(T)$ | $d_i(C)$ | $Y_i(T)$ | $Y_i(C)$ |
|---|----|-------|----------|----------|----------|----------|----------|----------|
| 1 | × | Т | × | ? | × | ? | × | ? |
| 2 | × | Т | × | ? | × | ? | × | ? |
| 3 | × | Т | × | ? | × | ? | × | ? |
| 4 | × | Т | × | ? | × | ? | × | ? |
| 5 | × | С | ? | × | ? | × | ? | × |
| 6 | × | С | ? | × | ? | × | ? | × |
| 7 | × | С | ? | × | ? | × | ? | × |
| 8 | × | С | ? | × | ? | × | ? | × |

• Individual Causal Effect: $E_i = Y_i(T) - Y_i(C)$

Donald B. Rubin

- Principal Stratum: $S_i = [D_i(T), D_i(C), d_i(T), d_i(C)];$ "Full" \Rightarrow strata considered property of patients.
- Principal Causal Effect: $\overline{E}_s = AVE_{i \in S}[Y_i(T) Y_i(C)]$. Average causal effect in principal stratum S.

rubin@stat.harvard.edu

Standard Assumptions

- Stable Unit Treatment Value Assumption (SUTVA):
 One patient's treatment assignment will not affect other patients' potential outcomes;
 No hidden versions of treatment and no hidden versions of control.
- Ignorable Treatment Assignment of T versus C: True for randomized experiment.

These are accepted by both EF and us.

Assumptions at the Individual Level

(1) Access Monotonicity

(1.A) General:
$$D_i(T) \geq D_i(C)$$
 and $d_i(C) \geq d_i(T)$

(1.B) Strong:
$$D_i(C) = 0$$
 and $d_i(T) = 0$

• (2) Side-Effect Monotonicity

(2.A) Negative:
$$D_i(T) \leq d_i(C)$$

(2.B) Positive:
$$D_i(T) \ge d_i(C)$$

(3) Exclusion Restriction:

$$D_i(T) = D_i(C), d_i(T) = d_i(C) \Rightarrow Y_i(T) = Y_i(C)$$

- (4) Perfect Blind: $D_i(T) = d_i(C)$
- (5) Equipercentile Equating of Compliances:

$$D_i(T) = F_D^{-1} \{ F_d[d_i(C)] \}$$

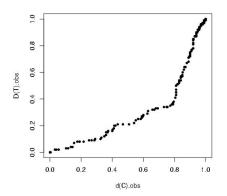
For LRC-CPPT:

- EF assumed: (1.B) and (5)

 ← true in expectation
- We assume: (1.B) and (2.A) ← weaker than (5)

EF's Assumption: Figure 3 Revisited

Figure: Q-Q Plot of Observed Drug and Placebo Compliances



Full Principal Stratification for LRC-CPPT

| i | Z_i | $D_i(T)$ | $D_i(C)$ | $d_i(T)$ | $d_i(C)$ | $Y_i(T)$ | $Y_i(C)$ |
|-------------|-------|----------|----------|----------|----------|----------|----------|
| 1 | Т | × | 0 | 0 | ? | × | ? |
| 2 | Т | × | 0 | 0 | ? | × | ? |
| | Т | × | 0 | 0 | ? | × | ? |
| n_T | Т | × | 0 | 0 | ? | × | ? |
| $n_{T} + 1$ | С | ? | 0 | 0 | × | ? | × |
| $n_{T} + 2$ | С | ? | 0 | 0 | × | ? | × |
| | С | ? | 0 | 0 | × | ? | × |
| n | С | ? | 0 | 0 | × | ? | × |

- Principal Stratum: $S_i = [D_i(T), 0, 0, d_i(C)] = [D_i, d_i]$ for notational simplicity.
- Recall, S_i is property of patients; modified later.
- Principal Causal Effect: $\overline{E}_s = AVE_{i \in S}[Y_i(T) Y_i(C)]$. Average causal effect in principal stratum S.

Parametric Model

Parametric model:

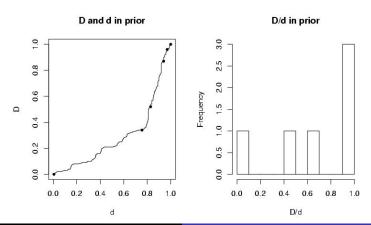
- $d_i | \theta, \rho \sim Beta(\alpha_1, \alpha_2);$ $\frac{D_i}{d_i} | d_i, \theta, \rho \sim Beta(\alpha_3, \alpha_4).$
- $Y_i(T)|D_i, d_i, \theta, \rho \sim N(\gamma_0 + \gamma_1 D_i + \gamma_2 D_i^2 + \gamma_3 d_i, \sigma_T^2).$
- $Y_i(C)|D_i, d_i, \theta, \rho \sim N(\beta_0 + \beta_1 D_i + \beta_2 d_i, \sigma_C^2)$.
- Partial correlation ρ : sensitivity parameter.

Therefore, $\theta = (\alpha_1, \alpha_2, \alpha_3, \alpha_4, \beta_0, \beta_1, \beta_2, \gamma_0, \gamma_1, \gamma_2, \gamma_3, \sigma_C, \sigma_T)$. **Prior distribution** $\pi(\theta|\rho)$:

- $\pi(\alpha_1, \alpha_2, \alpha_3, \alpha_4 | \rho)$: corresponds to adding 6 extra observations with complete (D, d) values.
- $\pi(\beta_0, \beta_1, \beta_2, \gamma_0, \gamma_1, \gamma_2, \gamma_3, \sigma_C, \sigma_T | \alpha_1, \alpha_2, \alpha_3, \alpha_4, \rho) \propto (\sigma_C \sigma_T)^{-2}$.

Parametric Model: Figure 4 - Prior Distribution

Figure: Prior Data Points for $\pi(\alpha_1, \alpha_2, \alpha_3, \alpha_4|\rho)$



Data Structure with Prior Data Points

| i | Z_i | $D_i(T)$ | $D_i(C)$ | $d_i(T)$ | $d_i(C)$ | $Y_i(T)$ | $Y_i(C)$ |
|-------------|-------|----------|----------|----------|----------|----------|----------|
| (1) | ? | × | 0 | 0 | × | ? | ? |
| (2) | ? | × | 0 | 0 | × | ? | ? |
| () | ? | × | 0 | 0 | × | ? | ? |
| (6) | ? | × | 0 | 0 | × | ? | ? |
| 1 | Т | × | 0 | 0 | ? | × | ? |
| 2 | Т | × | 0 | 0 | ? | × | ? |
| | Т | × | 0 | 0 | ? | × | ? |
| n_T | Т | × | 0 | 0 | ? | × | ? |
| $n_{T} + 1$ | С | ? | 0 | 0 | × | ? | × |
| $n_T + 2$ | С | ? | 0 | 0 | × | ? | × |
| | С | ? | 0 | 0 | × | ? | × |
| n | С | ? | 0 | 0 | × | ? | × |

Computation: MCMC

This is a missing data problem, and we can use MCMC to make Bayesian inference:

- Parameters: θ (fixed ρ)
- Key missing data: missing D_i or missing d_i
- In MCMC, given parameters θ , draw key missing data D_i or d_i ; then given missing data, draw parameters; iterate until convergence.
- After convergence, we can simulate missing $Y_i(T)$ or $Y_i(C)$ and obtain a set of complete data.

Therefore, we can get the posterior distribution of every estimand: θ , principal causal effects, principal stratum of each patient,...



Figure 5 - Scientific Estimands

Principal Causal Effects: Posterior Median and 95% Interval with $\rho=0$

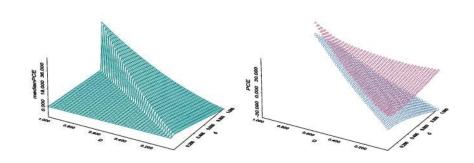
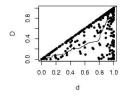
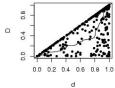
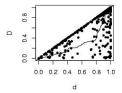


Figure 6 - Diagnostic Checks of EF's Assumption

Four Posterior Draws of Principal Strata for All the Patients with $\rho=0$







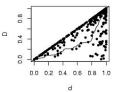
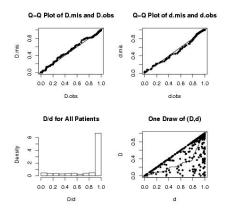


Figure 7 - Diagnostic Checks of Our Model

One Posterior Draw of D.mis and d.mis with $\rho = 0$



Sensitivity Analysis

Posterior Medians and Intervals of PCE for Different Values of ρ

| (D, d) | (1, 1) | (0.68, 0.89) | (0, 1) | (0, 0) |
|---------------|-------------|--------------|----------------|--------------|
| $\rho = -0.2$ | 49 (39, 59) | 24 (18, 30) | -10 (-40, 25) | 4 (-6, 14) |
| $\rho = 0$ | 50 (39, 59) | 24 (17, 30) | -13 (-42, 27) | 5 (-6, 16) |
| $\rho = 0.2$ | 50 (39, 59) | 23 (16, 29) | -11 (-47, 27) | 5 (-7, 18) |
| $\rho = 0.4$ | 50 (40, 59) | 23 (16, 29) | -6 (-43, 34) | 6 (-7, 20) |
| $\rho = 0.6$ | 51 (39, 62) | 22 (15, 30) | -10 (-43, 30) | 7 (-8, 23) |
| $\rho = 0.8$ | 52 (38, 63) | 22 (11, 33) | -8 (-62, 68) | 6 (-11, 28) |
| $\rho = 0.9$ | 51 (37 ,66) | 22 (6, 36) | -1 (-74 ,79) | 9 (-25 ,41) |

Understanding Principal Strata

Meaning of D_i and d_i :

- d_i: compliance to placebo indicates patient i's psychological compliance status.
- D_i: compliance to drug includes both patient i's psychological compliance status and his tolerance to negative side effects of the drug.
- d_i is more "fundamental" or "personal" than D_i .
- But D_i hints at possibility of estimating dose-response.
- Similar comments in EF.

Estimating the Dose-Response Relationship

Dose-Response within the Principal Stratification Framework:

- To estimate a dose-response relationship, we need a hypothetical experiment where different doses of drug are randomly assigned and strictly enforced (Also in EF).
- In the EF data, we need an additional assumption: for each cohort of patients with the same d, the assignment of D is stochastic and "latent ignorable" (Frangakis and Rubin 1999).
- With this additional assumption, we need a modified Principal Stratification framework, where D_i is no longer a stratum indicator.

Estimating the Dose-Response Relationship

Specific hypothetical experiment:

- Measure d_i^* = baseline compliance for each patient.
- Randomly divide patients into Treatment and Control.
- In the treatment group, stochastically assign dose $Z_{Di} \leq d_i^*$ according to a certain "rule".
- In the control group, assign full placebo and measure d_i .
- We notice $d_i = d_i^*$ in the control group, then "lose" d_i^* in the control group and in the treatment group.
 - \Rightarrow Non-ignorable assignment of Z_{Di} , ...but latent ignorable given d_i^* .
- Also, "forget" the rule for the assignment of Z_{Di} .



Estimating the Dose-Response Relationship

Modified Principal Stratification Framework for Dose-Response

| i | d_i^* | Z_i | Z_{Di} | $d_i(T)$ | $d_i(C)$ | $Y_i(T_0)$ | | $Y_i(T_D)$ | | $Y_i(T_1)$ | $Y_i(C)$ |
|-----------|---------|-------|----------|----------|----------|------------|---|------------|---|------------|----------|
| 1 | ? | Т | T_0 | 0 | ? | * | 2 | ? | ? | ? | ? |
| | ? | Т | | 0 | ? | ••• | | | | | ? |
| | ? | Т | T_D | 0 | ? | ? | ? | * | ? | ? | ? |
| | ? | Т | | 0 | ? | | | | | | ? |
| n_T | ? | Т | T_1 | 0 | ? | ? | ? | ? | ? | * | ? |
| $n_T + 1$ | ? | С | ? | 0 | * | ? | ? | ? | ? | ? | * |
| | ? | С | ? | 0 | * | ? | ? | ? | ? | ? | * |
| | ? | С | ? | 0 | * | ? | ? | ? | ? | ? | * |
| n | ? | С | ? | 0 | * | ? | ? | ? | ? | ? | * |



Modified Parametric Model

- $d_i|\theta, \rho \sim Beta(\alpha_1, \alpha_2);$ $\frac{Z_{Di}}{d_i}|d_i, \theta, \rho \sim Beta(\alpha_3, \alpha_4).$
- $Y_i(Z_{Di})|Z_{Di}, d_i, \theta, \rho \sim N(\gamma_0 + \gamma_1 Z_{Di} + \gamma_2 Z_{Di}^2 + \gamma_3 d_i, \sigma_T^2).$
- $Y_i(C)|d_i, \theta, \rho \sim N(\beta_0 + \beta_2 d_i, \sigma_C^2)$.
- Partial correlation ρ: sensitivity parameter.
- The prior distribution remains "the same" for (Z_{Di}, d_i) .
- The only difference is that the regression of $Y_i(C)$ cannot depend on randomly assigned (given d_i) dose, Z_{Di} , in treatment group.

Figure 8 - Dose-Response Results

Dose-Response and Sensitivity Analysis

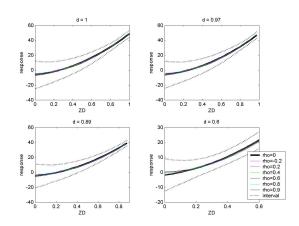
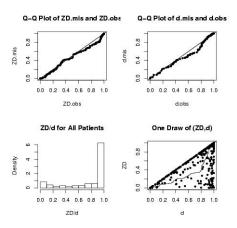


Figure 9 - Diagnostic Checks of Our Modified Model

One Posterior Draw of ZD.mis and d.mis with $\rho = 0$



Discussion of the Dose-Response Results

- Full Principal Stratification results are not causal for dose-response, but descriptive given principal strata.
- Dose-response results are causal under debatable assumptions.
- Is $Pr(Z_D|d, \{Y\}) = Pr(Z_D|d)$, "nature's randomization" of dose Z_{Di} given d_i , plausible?
- Or do we need Pr(Z_D|d, {Y}, M) = Pr(Z_D|d, M), where M refers to medical side effects of the drug beyond d?
 ⇒ Not latent ignorable given d, but latent ignorable given d and M.
- Sensitivity analysis to M? ⇒ future work.



Main References

- Efron, B., Feldman, D. (1991), "Compliance as an Explanatory Variable in Clinical Trials," *Journal of the* American Statistical Association 86, 9-17.
- Frangakis, C.E. and Rubin, D.B. (2002), "Principal Stratification in Causal Inference," Biometrics 58, 20-29.
- Hirano, K., Imbens, G.W., Rubin, D.B. and Zhou X.-H.
 (2000), "Assessing the Effect of an Influenza Vaccine in an Encouragement Design," *Biostatistics* 1, 69-88.
- Frangakis, C.E. and Rubin, D.B. (1999), "Addressing Complications of Intention-to-Treat Analysis in the Combined Presence of All-or-None Treatment -Noncompliance and Subsequent Missing Outcomes," *Biometrika* 86, 365-379.