# Employment protection, segmentation and wage inequality in Portugal

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## Motivation

What makes Portugal an interesting case to study wage inequality?

- Large supply shifts from 2.5% to 15% of college graduates.
- Institutional setting: Continental Europe; Segmentation.
- Open/integrated economy (technological changes; globalization).

# Part I Inequality: Aggregate analysis

## Overview

### **Previous studies**

Machado & Mata (2001), Martins & Pereira (2004), Cardoso (2007): Cover the earlier period (up to mid 90s) and tended to focus on college wage premium.

**Cardoso (1998), Machado & Mata (2005)**: for the overlapping period, inequality analysis is line with our findings.

## Upper-tail (90/50): 1982-2009

Levels:

Twice as much as in Germany;
 Starts at lower levels than the U.S., but ends up higher;

• Changes:

Portugal: +20 log points (1982-95); +11 l.p. (1996-09);
 United States & Germany: +4 and +5 l.p. (80s & 90s);

[Sources: Autor, Katz & Kearny, 2008 (United States); Schönberg, Dustmann & Ludsteck, 2009 (Germany)]

## Lower-tail (50/10): 1982-2009

- Levels:
  - United States > Portugal > Germany;
- Changes:
  - Portugal: 1982-95: 8 l.p.; 1996-09: Male/Female -6/3 l.p.
  - United States and Germany:
     1980s: 5/8 l.p. (M/F); 1990s: -1 l.p.;
     1980s: 3 l.p.; 1990s: 6 l.p.;

#### Market and institutional factors: 1982-1995

Increasing inequality:

- Low supply of skills;
  - 6 or less years of schooling: 69% in 1982, 44% in 1995;
  - college degree: 3% in 1982, 5% in 1995;
- No significant introduction of new institutions
- Skill-biased technological change is the main explanation.

#### It's a demand dominated story.

### Market and institutional factors: 1996-2009

Increasing upper-tail; but decreasing lower-tail inequality:

- Large shift: 5% to 15% of college graduates from 1995 to 2009;
- Real minimum wage increased in the late 1990s;
- Polarization;

#### It's a demand and supply story.

## Data

#### Administrative data: Quadros de Pessoal

- 1982-2009
- All (almost) salaried workers;
- Annual (October's spnashot)

## Analysis

#### Supply and demand

Methodological setting: Katz & Murphy (1992);

CES: 
$$Q_t = \left[\alpha_t \left(a_t N_{ct}\right)^{\rho} + (1 - \alpha_t) \left(b_t N_{nt}\right)^{\rho}\right]^{\frac{1}{\rho}}$$
 (1)  
 $ln\left(\frac{w_{ct}}{w_{nt}}\right) = \left(\frac{1}{\sigma}\right) \left[D_t - ln\left(\frac{N_{ct}}{N_{nt}}\right)\right],$  (2)

- Supply factors,  $\frac{N_{ct}}{N_{nt}}$  (college/non-college relative supply);
  - Elasticity of substitution,  $\sigma = \frac{1}{1-\rho}$ ;
- Demand factors (time trend), D<sub>t</sub>;
- Good 'Wage gap' model (fitted vs. actual);

#### College/Non-college: Elasticity of substitution



College/Non-college log wage gap										
	(1)	(2)	(3)	(4)	(5)					
College/Noncollege relative supply	-0.678	-0.683	-0.713	-0.481	-0.378					
	0.129	0.144	0.127	0.091	0.098					
Log real minimum wage			-0.610		-0.211					
			0.400		0.246					
Natural unemployment rate				-0.038	-0.043					
				0.006	0.007					
Time	0.052	0.052	0.059	0.043	0.045					
	0.008	0.009	0.010	0.006	0.006					
Time*1995		0.001			-0.004					
		0.003			0.001					
Constant	-1.109	-1.116	1.044	-0.409	0.632					
	0.349	0.381	1.450	0.258	0.869					
No. of observations	24	24	24	24	24					
$R^2$	0.777	0.777	0.800	0.911	0.937					

Katz & Murphy (1992): Fitted vs. Actual Wage Gap



### Experience levels: College/Non-college log wage gap

- Card & Lemieux (2001): Impact of experience
  - (same education, different experience): not perfect substitutes;
- Large supply shift concentrated in the 1990s;
  - Most noticeable in the younger cohort;
  - Visible impact on unconditional inequality;
- Split workers into 4 experience groups and compute the owngroup relative college/non-college supply and wage gaps;

#### 4 experience groups: C/NC wage gap

	Potential experience groups									
	All groups 0-9 years		10-19 years		20-29 years		30-39 years			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
Own minus aggregate supply	-0.324	-0.323	-0.209	-0.254	-0.584	-0.548	0.132	0.059	0.427	0.200
	0.010	0.009	0.129	0.110	0.043	0.112	0.161	0.064	0.197	0.116
Aggregate supply	-0.578	-0.400	-0.614	-0.435	-0.459	-0.464	-0.242	-0.106	-0.163	-0.034
	0.105	0.107	0.187	0.176	0.064	0.068	0.308	0.125	0.085	0.054
Log real minimum wage		-0.174		-0.188		-0.016		-0.141		-0.130
		0.314		0.325		0.189		0.198		0.160
Natural unemployment rate		-0.036		-0.027		-0.004		-0.052		-0.027
		0.008		0.008		0.012		0.005		0.004
Time	0.048	0.042	0.044	0.038	0.040	0.041	0.035	0.033	0.040	0.030
	0.007	0.008	0.011	0.011	0.004	0.005	0.017	0.008	0.007	0.005
Constant	-0.734	0.543	-0.782	0.544	-0.325	-0.267	0.194	1.291	0.630	1.451
	0.284	1.107	0.585	1.109	0.168	0.665	0.847	0.731	0.260	0.562
No. of observations	96	96	24	24	24	24	24	24	24	24
$R^2$	0.882	0.906	0.738	0.848	0.972	0.972	0.903	0.987	0.969	0.991

Own and aggregate supply: only significant for the two less experienced groups.





21 I.p. increase in the difference between the C/NC wage gap of the less- and most-experienced. Using the

Using the estimated own-group elasticity: Explains 17 I.p.; 82% of total difference;

#### Minimum wage

	M	ale	Ferr	ale	
	90/50	50/10	90/50	50/10	
College/Noncollege relative supply	-0.065	-0.185	-0.267	0.006	
	0.037	0.048	0.072	0.033	
Log real minimum wage	-0.030	-0.265	-0.331	-0.457	
	0.110	0.141	0.210	0.098	
Natural unemployment rate	-0.011	-0.016	-0.010	-0.008	
	0.003	0.004	0.005	0.003	
Time	0.020	0.018	0.035	0.007	
	0.003	0.003	0.005	0.002	
Constant	0.528	0.901	1.085	1.924	
	0.387	0.499	0.743	0.346	
No. of observations	24	24	24	24	
$R^2$	0.992	0.828	0.969	0.934	

Katz-Murphy: 90/50 and 50/10 wage percentiles ratios as dependent variables;

Females 50/10 ratio is "explainable" with minimum wage developments.

#### Counterfactuals analysis: DiNardo, Fortin & Lemieux (1996)

Let f(w|T = t) be the **observed wage density** at time t. It can be decomposed into the density of observed wage **conditional on attributes** x at time t, g(w|x, T = t) and the density of the same **attributes**, h(x|T = t).

$$f(w|T = t) = \int g(w|x, T = t)h(x|T = t)dx$$

**Counterfactual**: wage distribution of t that would have prevailed if attributes were those of year t'.

#### **Overall inequality**



Price effect: Vertical distance bewteen curves; Composition effect: Movement along curves

Price effect in 82-95: 24.5  $\Rightarrow$  13.2 (1995's X); Composition effect in 96-09: 15.2  $\Rightarrow$  -7.1 (1995's X); Price effect in 82-95:  $8.5 \Rightarrow 6.1 \ (1995's \ X);$ Price effect in 96-09:  $-5.6 \Rightarrow -11.1 \ (1995's \ X);$ 



Price effect: Vertical distance bewteen curves; Composition effect: Movement along curves

#### Polarization

Decrease in 50/10 wage inequality. Explanations?

- SBTC cannot account for the decrease;
- Minimum wage helps marginally;
- But other demand factors may be at play (Goos & Manning, 2007); we can gauge them by:
  - Employment shares by occupational skill;
  - Real wage variation by wage percentile;



#### Polarization: Real wage log variation

# Part II Inequality: Causal analysis

## Labor Code 2004 reform: A quasi-experiment

**Fair dismissals**: Firing a worker implies: (i) written procedures; (ii) witnesses interviews involving the works council.

**New law:** firms with **11 to 20** workers have to comply with this **additional procedural requirements**. Before 2004, only firms **21**+ workers.

- Treatment firms: 11-20 workers
- Control firms: 21-50 workers
- Before: 2002-2003
- After: 2004-2008

## Data

Quadros de Pessoal: 2002 - 2008

Our analysis starts in **2002**, the first year for which the information on the **type of contract is available**, and ends in 2008, to avoid the influence of the 2009 Labor Code revision.

Common t	trend
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		90	/50		50/10				
	Base	wage	Total	wage	Base	wage	Total wage		
	Monthly	Hourly	Monthly	Hourly	Monthly	Hourly	Monthly	Hourly	
$Treat\timesTime$	-0.163	-0.643	-0.113	-0.285	0.007	0.076	0.027	0.074	
	(0.041)	(0.220)	(0.187)	(0.573)	(0.878)	(0.835)	(0.632)	(0.848)	
Treat	0.338	1.743	0.291	1.377	0.062	0.991	0.009	0.591	
	(0.084)	(0.176)	(0.158)	(0.257)	(0.585)	(0.251)	(0.945)	(0.528)	
Time	0.116	-0.245	0.090	-0.356	-0.036	-1.013	-0.055	-1.000	
	(0.083)	(0.574)	(0.198)	(0.383)	(0.359)	(0.001)	(0.228)	(0.001)	
No of observations	53278								

#### Match (worker $\times$ firm) fixed-effects with clustering.

Control variables: (i) number of workers as a proxy for firm size; (ii) firm age (indicator variables: 1, 2, ..., 10, 11-15, 16-20, and more than 20 years); (iii) sector; (iv) region, (v) foreign ownership majority; (vi) gender; (vii) nationality; (viii) age; (ix) education; (x) white and blue collar; (xi) workers on a (regulated) minimum wage; (xii) tenure.

#### Quasi-experimental evidence: Impact on inequality

		90	/50		50/10				
	Base	Base wage		wage	Base wage		Total wage		Main messages:
	Monthly	Hourly	Monthly	Hourly	Monthly	Hourly	Monthly	Hourly	0
All contracts	0.063	0.727	0.043	0.368	0.061	0.442	0.065	0.572	Adjustment in
	(0.144)	(0.007)	(0.309)	(0.129)	(0.018)	(0.019)	(0.024)	(0.002)	hase wages
				211	369				buse wages
Open-ended contracts	-0.003	0.390	-0.055	-0.057	0.071	0.504	0.094	0.613	Primarily on
	(0.948)	(0.193)	(0.244)	(0.831)	(0.020)	(0.019)	(0.004)	(0.003)	(new) OEC.
	201243								
Fixed-term contracts	0.028 (0.688)	1.113 (0.013)	0.051 (0.472)	1.037 (0.012)	0.014 (0.765)	0.239 (0.472)	-0.017 (0.737)	0.106 (0.745)	Inequality among FTC unaffected.
				143	201				

## Conclusion

- Inequality increased from 1982 to 2006
- Supply & demand have been the "designated drivers"
- Institutional settings took the "passenger seat";

Novelty of our results: Causal relationship between employment protection and inequality. However, this is still work in progress.

# Thank you.